

考試科目	數理統計	所別	統計所	考試時間	5月24日 星期六	第一節
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1. (15 分) For uniform distribution $U(0,1)$ random variables U_1, U_2, \dots , define

$N = \min\{n : \sum_{i=1}^n U_i > 1\}$. That is, N is the number of random numbers that must be summed to exceed 1. Compute the density function of N , $E(N)$, and $\text{Var}(N)$.

2. (20 分) Let X and Y be independently and identically distributed $N(0, \sigma^2)$ random variables. Show that:

(a) $X^2 + Y^2$ and $\frac{X}{\sqrt{X^2 + Y^2}}$ are independent.

(b) θ is uniformly distributed on $\left(-\frac{\pi}{2}, \frac{\pi}{2}\right)$, if $\theta = \sin^{-1} \frac{X}{\sqrt{X^2 + Y^2}}$.

(c) X/Y has a Cauchy distribution.

3. (15 分) Sequential Probability Ratio Test (SPRT), developed by Wald, can be used to test simple null hypothesis (H_0) versus simple alternative hypothesis (H_1). The test is based on the ratio of likelihood function under H_0 and H_1 , i.e.,

$$\lambda_n = \frac{\prod_{i=1}^n f(X_i | H_0)}{\prod_{i=1}^n f(X_i | H_1)} \quad \text{or} \quad \log \lambda_n = \sum_{i=1}^n \log \left(\frac{f(X_i | H_0)}{f(X_i | H_1)} \right),$$

where X_1, \dots, X_n are a random sample from p.d.f. $f(x | H_J)$ with $J = 0$ or 1 . The rule of applying SPRT is to continue sampling until $\log \lambda_n \geq B$ or $\log \lambda_n \leq A$,

where $B > A$. The values of A and B depend on the type I error α and type II

error β . It is proved that the SPRT minimizes the sample size under H_0 and

H_1 , given the error probabilities α and β . Suppose the sample is drawn from

Bernoulli distribution $B(1, \theta)$ with $H_0: \theta = 3/8$ and $H_1: \theta = 1/2$. Let $\alpha = \beta = 0.05$.

Compute the expected sample size if H_0 is true, and the expected sample if H_1 is true.

備 考 | 試題隨卷繳交

命題委員： (簽章) 年 月 日

國立政治大學九十七 學年度研究所**博士**班入學考試命題紙 第之頁，共之頁

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4. (15 分) Let X_1, \dots, X_n be a random sample from Bernoulli(p) distribution and

$$Y_n = \frac{\sum_{i=1}^n X_i}{n}. \quad (\text{a}) \text{ Show that } \sqrt{n}(Y_n - p) \rightarrow N(0, p(1-p)) \text{ in distribution.}$$

$$(\text{b}) \text{ Show that for } p = 1/2, n[Y_n(1-Y_n) - \frac{1}{4}] \rightarrow -\frac{1}{4}\chi_1^2 \text{ in distribution.}$$

5. (15 分) Let X_1, \dots, X_n be a random sample from uniform distribution

$$U(\theta - \frac{1}{2}, \theta + \frac{1}{2}). \text{ Show that for any } T \text{ such that } X_{(n)} - \frac{1}{2} \leq T \leq X_{(1)} + \frac{1}{2} \text{ is a}$$

maximum likelihood estimate of θ , where $X_{(1)} = \min\{X_1, \dots, X_n\}$ and

$$X_{(n)} = \max\{X_1, \dots, X_n\}.$$

6. (20 分) Let X_1, \dots, X_n be a random sample from uniform distribution $U(0, \theta)$ and

$$R = X_{(n)} - X_{(1)}, \text{ where } X_{(n)} = \max\{X_1, \dots, X_n\} \text{ and } X_{(1)} = \min\{X_1, \dots, X_n\}.$$

(a) Find the distribution of R .

(b) Show that a confidence interval for θ , based on R , with confidence coefficient $1-\alpha$ is of the form $[R, R/c]$, where c is the root of the equation $c^{n-1}[n - (n-1)c] = \alpha$.

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