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# Adaptive sampling interval cause-selecting control charts 

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#### Abstract

This article considers the statistical adaptive process control for two dependent process steps. We construct an adaptive sampling interval $Z_{X}$ control chart to monitor the quality variable produced by the first process step, and use the adaptive sampling interval $Z_{e}$ control chart to monitor the specific quality variable produced by the second process step. By using the proposed adaptive sampling interval control charts, we can quickly detect and distinguish which process step is out of control. The performance of the proposed adaptive sampling interval control charts is measured by the adjusted average time to signal (AATS), which was derived by a Markov chain approach, for an out-of-control process. An empirical automobile braking system example shows the application and the performance of the proposed adaptive sampling control charts in detecting shifts in process means. Some numerical results obtained demonstrated that the performance of the proposed adaptive sampling cause-selecting control charts outperforms the fixed sampling interval cause-selecting control charts.


Keywords Adaptive sampling interval • Control charts • Dependent process steps • Adjusted average time to signal

## Introduction

Control charts are important tools in statistical quality control. They are used to effectively monitor and determine whether a process is in control or out of control. Shewhart [20] first developed the $\bar{X}$ control chart to monitor the process mean. The control chart is easy to implement and, hence, it has been widely used for industrial process control. Usually, Shewhart $\bar{X}$ control charts monitor a process by taking equal samples of size at a fixed sampling interval.

[^0]Once the sample statistic falls outside of its control limits, it indicates that the process is out of control, and that it should be stopped and repaired. Shewhart $\bar{X}$ control charts provide good performance in detecting large shifts in the process mean. However, small and moderate shifts in process means are expected in reality. There have been several alternatives proposed to improve this problem in recent years. Some of the alternatives include Shewhart $\bar{X}$ control charts combined with cumulative sum (CUSUM) control charts [10], Shewhart $\bar{X}$ control charts with run rules [9], exponentially weighted moving average (EWMA) control charts [11], and Shewhart $\bar{X}$ control charts with adaptive sampling interval or/and sample size [16], [13].

The idea of adaptive sampling interval (ASI) Shewhart $\bar{X}$ control charts is that the sampling interval should be short if the previous sample shows some indication of a shift in process mean, and long if there is no indication of a shift. Reynolds et al. [16] propose Shewhart $\bar{X}$ control charts with an adaptive sampling interval where the time period between samples is varied, depending on the location of the previous sample statistic on the charts. Their paper demonstrates that dual sampling intervals are better, and that the proposed control chart can detect small and moderate shifts in process mean faster than a fixed sampling interval (FSI) Shewhart $\bar{X}$ control chart. An ASI Shewhart $\bar{X}$ control chart with run rules is introduced by Cui and Reynolds [6]. The control chart signals not only a sample statistic falling outside the control limits, but also a run of a specified length occurred in a specified region. It offers better performance than ASI Shewhart $\bar{X}$ control charts without run rules and FSI Shewhart $\bar{X}$ control charts with or without run rules. Reynolds et al. [15] propose the CUSUM control chart with adaptive sampling interval to monitor the shift in process mean by extending the approach Reynolds et al. [16]. Their paper shows that the ASI CUSUM control chart is substantially more efficient than either the ASI Shewhart $\bar{X}$ control chart or the FSI CUSUM control chart for small and moderate shifts in process mean. Runger and Pignatiello [18] propose a one-
sided control limit for an adaptive ASI Shewhart $\bar{X}$ control chart, and show that the improvement in detecting the process mean is better than a two-sided adaptive ASI Shewhart $\bar{X}$ control chart. It also shows that both one-sided and two-sided adaptive Shewhart $\bar{X}$ control charts detect small and moderate shifts in the process mean faster than an FSI Shewhart $\bar{X}$ control chart. Saccucci et al. [19] introduce the properties and performance of a EWMA control chart with adaptive sampling interval. Their paper demonstrates that the ASI EWMA control chart always outperforms the FSI EWMA control chart, as well as the ASI Shewhart $\bar{X}$ control chart, especially in detecting small and moderate shifts. Amin and Miller [1] propose a robust ASI $\bar{X}$ control chart. The chart is constructed based on a non-normal distributed quality characteristic. Reynolds [14] proposes a varied sampling interval $\bar{X}$ control chart with sampling at fixed times (VSIFT), and shows that the performance of the VSIFT $\bar{X}$ control chart is better than the FSI Shewhart $\bar{X}$ control chart and is relatively easy to set up and implement in practice.

Similar to ASI control charts, some papers consider adaptive sample size (ASS) control charts [2, 7, 13, 24]. Some authors combine adaptive sample size and sampling interval (ASSI) control charts [3, 12, 17]. Costa [5] proposes the adaptive parameter (AP) $\bar{X}$ control chart, and also discusses adaptive parameter (AP) $\bar{X}$ and $R$ control charts [4]. Epprecht et al. [8] proposed adaptive control charts for attributes. A survey of recent developments in the adaptive control charts sees Tagaras [21].

All the above papers only consider a single process step. However, most recent products are produced from several different dependent process steps. In multiple process steps, a Shewhart control chart is often used to monitor the process state of each individual step. If the process steps are independent, then, using a Shewhart control chart to monitor each individual step is meaningful. However, many process steps are not independent, and, thus, the control charts are difficult to interpret the correct process state. An alternative approach is to use a multivariate control chart, such as a Hotelling $T^{2}$ control chart, to monitor dependent process steps. The disadvantages of using a Hotelling $T^{2}$ control chart are that it assumes that all quality characteristics in the chart are multivariate normal random variables, and there is difficulty in interpretation for which process step is out of control when there is a signal from the control chart. For the two dependent process steps, a powerful and popular approach proposed by Zhang [23] is using cause-selecting control charts. The cause-selecting control chart is constructed for an outgoing quality variable only after the observations have been adjusted for the effect of an incoming quality variable. The advantage of this approach is that, once there is a signal, it is easy to determine if the second step of the process is out of control. Wade and Woodall [22] review and analyze the cause-selecting control chart and examine the relationship between the cause-selecting control chart and the Hotelling $T^{2}$ control chart. In their opinion, the cause-selecting control chart outperforms the Hotelling $T^{2}$ control chart.

Yang [25] proposed economic control charts for two dependent process steps. Yang [26] considers a statistical process control for two dependent process steps with a failure mechanism. Yang and Chen [28] propose two dependent processes control for two failure mechanisms from an economic viewpoint. Yang [27] addresses dependent processes control for over-adjusted process means. Yang and Yang [29] study the effects of imprecise measurement on controlling two dependent process steps for autocorrelated observations. However, using adaptive control charts to monitor and distinguish which process step is out of control has not been addressed.

The purpose of this paper is to study the ASI $Z_{X}$ and $Z_{e}$ control charts to monitor the process means on the dependent first process step and the second process step, respectively. Following Costa [3], the shifts in the process means do not occur at the beginning, but instead at some random time in the future, and the occurrence time of the shift is assumed to be an exponentially distributed random variable. The performance of the ASI $Z_{X}$ and $Z_{e}$ control charts is measured by the Adjusted average time to signal (AATS). An empirical example is given to illustrate the application and performance of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts. Finally, some numerical comparison results show that the performance of the proposed ASI charts outperforms the FSI $Z_{X}$ and $Z_{e}$ control charts.

## Process description

Throughout this article, a two-step process is assumed. Denote $X$ to be the quality measurement of interest for the first step and $Y$ to be the quality measurement of interest for the second step. The two steps of the process are dependent, and the second step is affected by the first step. Suppose that the values of $X$ cannot be observed in the first step, but can be measured at the end of the second step. Hence, a sample of size 1 is taken at a varied sampling interval, and the paired observations $\left(X_{i}, Y_{i}\right)$ are measured at the end of the second step. It is assumed that the samples are mutually independent. Figure 1 shows a two-step process and the interested quality variables $(X, Y)$ for the process.

The quality variable $X$ is assumed to follow a normal distribution with mean $\mu_{X}$ and a constant standard deviation $\sigma_{X}$ when the first step is in control. The quality variable $Y$ is affected by $X$ and the relationship between $Y$ and $X$ is expressed as:
$Y_{i} \mid X_{i}=f\left(X_{i}\right)+\varepsilon_{i}, \quad i=1,2,3, \ldots, n$.


Fig. 1 A two-step process

The variable $\varepsilon_{i}$ is the random error and $\varepsilon_{i} \sim \operatorname{NID}\left(0, \sigma_{\varepsilon}^{2}\right)$ when the second step is in control. To monitor the two dependent process steps effectively, two control charts are constructed to control the first step and the second step respectively. To monitor the first step, the individual $X$ control chart is set up based on the in control distribution of $X$. To monitor the second step, the specific quality of the second step is specified by adjusting the effect of $X$ on $Y$; that is, the specific quality is presented by the cause-selecting values, $e_{i}=Y_{i}\left|X_{i}-\widehat{Y}_{i}\right| X_{i}$. The cause-selecting control chart is set up based on the in-control distribution of causeselecting values. Hence, $X \sim \mathrm{~N}\left(\mu_{X}, \sigma_{X}^{2}\right)$ and $e \sim \mathrm{~N}\left(0, \sigma_{\varepsilon}^{2}\right)$ when the first step and the second step are both in control.

We assume the process is in control at the beginning, and two assignable causes, say AC1 and AC2, may occur at the process randomly. AC1 may only occur at the first step and influences a shift in the distribution of X to $\mathrm{X} \sim \mathrm{N}\left(\mu_{X}+\right.$ $\delta_{1} \sigma_{X}, \sigma_{X}^{2}$ ). AC2 may only occur at the second step and influences a shift in the distribution of $e$ to $e \sim \mathrm{~N}\left(\delta_{2} \sigma_{\varepsilon}, \sigma_{\varepsilon}^{2}\right)$. When both of AC1 and AC2 occur, the process mean of $X$ shifts from $\mu_{X}$ to $\mu_{X}+\delta_{1} \sigma_{X}$, and the process mean of $e$ shifts from 0 to $\delta_{2} \sigma_{\varepsilon}$. The assignable cause occurs according to an exponential distribution with parameter $\lambda_{i}, i=1,2$. Table 1 shows possible distributions of $X$ and $e$.

## Principles of the ASI $Z_{X}$ and $Z_{e}$ control charts for two dependent process steps

To control and diagnose the two dependent process steps effectively, the adaptive sampling interval $X$ and causeselecting control charts are constructed based on the incontrol $X$ and $e$ distributions. The sampling interval is varied and it depends on the locations of the previous sample statistics. Usually, there are three lines (upper control limit, central limit, and lower control limit) for a control chart. Here, we add a warning threshold on each of the proposed control charts. Hence, there are two symmetric warning lines between the control limits. Consequently, the structures of the ASI $X$ and cause-selecting control charts are as shown in Fig. 2.

To monitor the two dependent process steps, a sample of size 1 is taken at the end of the second step, the paired observations $\left(X_{i}, Y_{i}\right)$ are measured, and the relationship between $Y$ and $X$ is determined. Then, the sample statistics $\left(X_{i}, e_{i}\right)$ are calculated and plotted on the ASI $X$ control chart and cause-selecting control chart, respectively. If the sample statistic falls between the warning lines, it is rea-

Table 1 Possible distributions of $X$ and $e$

| $X$ |  | $e$ |  |  |
| :--- | :--- | :--- | :--- | :--- |
| Does AC1 occur? | Mean of $X$ |  | Does AC2 occur? | Mean of $e$ |
| No | $\mu_{X}$ | No | 0 |  |
| Yes | $\mu_{X}+\delta_{1} \sigma_{X}$ | No | 0 |  |
| No | $\mu_{X}$ | Yes | $\delta_{2} \sigma_{\varepsilon}$ |  |
| Yes | $\mu_{X}+\delta_{1} \sigma_{X}$ | Yes | $\delta_{2} \sigma_{\varepsilon}$ |  |

$$
\begin{aligned}
& \mathrm{UCL}_{\mathrm{X}}=\mu_{X}+k_{X} \sigma_{\mathrm{X}} \\
& \mathrm{UWL}_{\mathrm{X}}=\mu_{X}+w_{X} \sigma_{\mathrm{X}} \\
& \mathrm{CL}_{\mathrm{X}}=\mu_{X} \\
& \mathrm{UWL}_{e}=k_{e} \sigma_{\varepsilon} \\
&=\mu_{X}-w_{X} \sigma_{\mathrm{X}} \\
& \mathrm{UWL}_{e}=w_{e} \sigma_{\varepsilon} \\
& \mathrm{LCL}_{\mathrm{X}}=\mu_{X}-k_{X} \sigma_{\mathrm{X}} \\
& \text { ASI X control chart } \mathrm{CL}_{e}=0 \\
& \mathrm{LWL}_{e}=-w_{e} \sigma_{\varepsilon} \\
& \mathrm{LCL}_{e}=-k_{e} \sigma_{\varepsilon}
\end{aligned}
$$

Fig. 2 The ASI $X$ and cause-selecting control charts
sonable to wait more time to take the next sample because there is no evidence that the process needs adjustment. If the sample statistic falls close to the control limits but not outside them, it is reasonable to wait less time to take the next sample because the process can be in need of adjustment. When at least one sample statistic falls outside of the control limits, we stop monitoring the process and start to repair the assignable cause.

To easily use the proposed charts for process engineers, we standardize sample statistics $X$ and $e$ as follows. Let the following:

$$
\begin{equation*}
Z_{X}=\frac{X-\mu_{X}}{\sigma_{X}} \tag{2}
\end{equation*}
$$

where $Z_{X} \sim \mathrm{~N}(0,1)$ when the first step is in control. Let the following:

$$
\begin{equation*}
Z_{e}=\frac{e}{\sigma_{\varepsilon}} \tag{3}
\end{equation*}
$$

where $Z_{e} \sim N(0,1)$ when the second step is in control.
Based on the in-control distributions of $Z_{X}$ and $Z_{e}$, the structures of the ASI $Z_{X}$ and $Z_{e}$ control charts are as shown in Fig. 3.

Before using the proposed ASI $Z_{X}$ and $Z_{e}$ control charts, we determine three sampling intervals, $t_{1}, t_{2}$, and $t_{3}$, where $t_{3}>t_{2}>t_{1}$. If both sample statistics $\left(Z_{X}, Z_{e}\right)$ fall between the warning limits, it is reasonable to choose a longer sampling interval, $t_{3}$, to take the next sample, as no evidence that the process needs adjustment. If both sample statistics fall outside the warning limits but inside the control limits, it is reasonable to choose the shorter sampling interval, $t_{1}$, to take the next sample, as the process can be in need of adjustment. If one of the sample statistics falls within the

$$
\begin{array}{rlrl}
\mathrm{UCL}_{z_{x}} & =k_{X} & \mathrm{UCL}_{z_{e}} & =k_{e} \\
U W L_{z_{x}} & =w_{X} & \mathrm{UWL}_{z_{e}} & =w_{e} \\
\mathrm{CL}_{\mathrm{z}_{\mathrm{x}}} & =0 & \mathrm{CL}_{\mathrm{z}_{\mathrm{e}}} & =0 \\
\mathrm{LWL}_{\mathrm{z}_{\mathrm{x}}} & =-w_{X} & \mathrm{LWL}_{z_{e}} & =-w_{e} \\
\mathrm{LCL}_{\mathrm{z}_{\mathrm{X}}} & =-k_{X} & \mathrm{LCL}_{\mathrm{z}_{\mathrm{e}}} & =-k_{e} \\
\text { ASI } & Z_{X} \text { chart } & \text { ASI } Z_{e} \text { chart }
\end{array}
$$

Fig. 3 The ASI $Z_{X}$ and $Z_{e}$ control charts
warning limit and the other falls outside the warning limit but inside the control limit, it is reasonable to choose a middle sampling interval, $t_{2}$. Once one sample statistic falls outside of its control limits, then the sampling stops and the process restarts. The first sampling interval is randomly chosen from $t_{1}, t_{2}$, and $t_{3}$, and the following sampling interval is dependent upon the locations of the previous sample statistics.

To express the relationship between the sampling interval, $t_{i}$, and the location of the sample statistics on the proposed charts, the ASI $Z_{X}$ and $Z_{e}$ control charts are divided into the following four regions:

$$
\begin{array}{rlrl}
I_{X 1}=\left[-w_{X}, w_{X}\right] & I_{e 1}= & {\left[-w_{e}, w_{e}\right]} \\
I_{X 2}=\left(-k_{X},-w_{X}\right) & I_{e 2}=\left(-k_{e},-w_{e}\right) \\
& \cup\left(w_{X}, k_{X}\right) & & \cup\left(w_{e}, k_{e}\right) \\
I_{X 3}= & \left(-k_{X}, k_{X}\right) & I_{e 3}=\left(-k_{e}, k_{e}\right) \\
I_{X 4}=\left(-\infty,-k_{X}\right] & I_{e 4}= & \left(-\infty,-k_{e}\right] \\
& \cup\left[k_{X}, \infty\right) & & \cup\left[k_{e}, \infty\right)
\end{array}
$$

According to the adaptive scheme, if the previous sample statistics $Z_{X}$ falls in region $I_{X 1}$ and $Z_{e}$ falls in region $I_{e 1}$, then we choose the longer sampling interval $t_{3}$. If one of the previous sample statistics falls in region $I_{X 1}$ or $I_{e 1}$, and the other falls in region $I_{X 2}$ or $I_{e 2}$, then we choose the sampling interval $t_{2}$. If both of the previous sample statistics fall in region $I_{X 2}$ and $I_{e 2}$, then we choose the shorter sampling interval $t_{1}$. If at least one of the sample statistics falls outside of the control limits, then we stop sampling. Their relationship is formulated as in Eq. 4:
$t_{k}=\left\{\begin{array}{l}t_{3} \text { if } Z_{X} \in I_{X 1} \cap Z_{e} \in I_{e 1} \\ t_{2} \text { if } Z_{X} \in I_{X 1} \cap Z_{e} \in I_{e 2} \\ t_{2} \text { if } Z_{X} \in I_{X 2} \cap Z_{e} \in I_{e 1} \\ t_{1} \text { if } Z_{X} \in I_{X 2} \cap Z_{e} \in I_{e 2}\end{array}, \boldsymbol{k}=1,2,3, \ldots\right.$.

No matter if the signal is a true or a false on the proposed $Z_{X}$ control chart, it indicates that AC1 may have occurred at the first step. Then, the process is stopped for repairs until the process is back in control, similar to the proposed $Z_{e}$ control chart.

## Determination of the warning limits on the ASI $Z_{X}$ and $Z_{e}$ control charts

The design parameters of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts are the sampling intervals $\left(t_{1}, t_{2}, t_{3}\right)$, the control limits $\left(k_{X}, k_{e}\right)$, and the warning limits $\left(w_{X}, w_{e}\right)$. The three sampling intervals are always determined by the process engineers, and the control limits are always fixed. This leaves the warning limits to be determined. In order to compare the performance of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts and the FSI $Z_{X}$ and $Z_{e}$ control charts after the process shifts, we should let the ASI $Z_{X}$ and $Z_{e}$ control charts and FSI $Z_{X}$
and $Z_{e}$ control charts have the same expected in-control process time. Hence, under the in-control process, we let the expected sampling time interval $E\left(t_{i}\right)$ of the proposed ASI control charts be equal to the sampling time interval of the FSI $Z_{X}$ and $Z_{e}$ control charts.

That is:

$$
E\left[t_{k} \mid \delta_{1}=0, \delta_{2}=0\right]
$$

$=t_{0} P\left(\right.$ plotted points within control limits $\left.\mid \delta_{1}=0, \delta_{2}=0\right)$,
where $t_{0}$ is the sampling time interval of the FSI $Z_{X}$ and $Z_{e}$ control charts. We have to note that the only difference between the FSI and ASI $Z_{X}$ and $Z_{e}$ control charts is that the former has no warning limits. Using the constraint in Eq. 5, we can derive the warning limits of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts. We assume that, once false alarm occurs, then the process stops sampling and restarts. Hence, the constraint can be expressed as follows:

$$
\begin{align*}
& t_{3} \times P\left(Z_{X} \in I_{X 1} \mid \delta_{1}=0, \delta_{2}=0\right) \\
& \quad \times P\left(Z_{e} \in I_{e 1} \mid \delta_{1}=0, \delta_{2}=0\right)+ \\
& \quad t_{2} \times P\left(Z_{X} \in I_{X 1} \mid \delta_{1}=0, \delta_{2}=0\right) \\
& \quad \times P\left(Z_{e} \in I_{e 2} \mid \delta_{1}=0, \delta_{2}=0\right)+ \\
& \quad t_{2} \times P\left(Z_{X} \in I_{X 2} \mid \delta_{1}=0, \delta_{2}=0\right) \\
& \quad \times P\left(Z_{e} \in I_{e 1} \mid \delta_{1}=0, \delta_{2}=0\right) \\
& \quad t_{1} \times P\left(Z_{X} \in I_{X 2} \mid \delta_{1}=0, \delta_{2}=0\right)  \tag{6}\\
& \quad \times P\left(Z_{e} \in I_{e 2} \mid \delta_{1}=0, \delta_{2}=0\right) \\
& \quad \times P(\text { at least one false alarm }) \\
& \quad=t_{0} \times P\left(-k_{x}<Z_{X, i-1}<k_{x} \mid \delta_{1}=0, \delta_{2}=0\right) \\
& \quad \times P\left(-k_{e}<Z_{e}<k_{e} \mid \delta_{1}=0, \delta_{2}=0\right) \\
& \quad \times P(\text { at least one false alarm })
\end{align*}
$$

Simplifying Eq. 6 gives:

$$
\begin{align*}
& 4 \Phi\left(w_{X}\right) \Phi\left(w_{e}\right)\left[t_{3}-2 t_{2}+t_{1}\right] \\
& \quad+2 \Phi\left(w_{X}\right)\left[-t_{3}+2 t_{2} \Phi\left(k_{e}\right)+t_{2}-2 t_{1} \Phi\left(k_{e}\right)\right] \\
& \quad+2 \Phi\left(w_{e}\right)\left[-t_{3}+2 t_{2} \Phi\left(k_{X}\right)+t_{2}-2 t_{1} \Phi\left(k_{X}\right)\right]  \tag{7}\\
& =t_{0}\left(2 \Phi\left(k_{X}\right)-1\right)\left(2 \Phi\left(k_{e}\right)-1\right)-t_{3}+2 t_{2} \Phi\left(k_{e}\right) \\
& \quad+2 t_{2} \Phi\left(k_{X}\right)-4 t_{1} \Phi\left(k_{X}\right) \Phi\left(k_{e}\right)
\end{align*}
$$

where $\Phi(\cdot)$ denotes the standard normal cumulative function.
Let $k_{X}=k_{e}=k$, Eq. 7 is simplified as:

$$
\begin{align*}
& 4 \Phi\left(w_{X}\right) \Phi\left(w_{e}\right)\left[t_{3}-2 t_{2}+t_{1}\right]+2\left[\Phi\left(w_{X}\right)+\Phi\left(w_{e}\right)\right] \\
& \quad \times\left[-t_{3}+2 t_{2} \Phi(k)+t_{2}-2 t_{1} \Phi(k)\right] \\
&= t_{0}(2 \Phi(k)-1)^{2}-t_{3}+4 t_{2} \Phi(k)-4 t_{1}(\Phi(k))^{2} \tag{8}
\end{align*}
$$

We can solve Eq. 8 by letting $w_{X}=w_{e}=w$, and it reduces the complication of implementing the proposed ASI $Z_{X}$ and $Z_{e}$ control charts.

Consequently:

$$
\begin{align*}
& 4 \Phi(w)^{2}\left[t_{3}-2 t_{2}+t_{1}\right] \\
& \quad+4 \Phi(w)\left[-t_{3}+2 t_{2} \Phi(k)+t_{2}-2 t_{1} \Phi(k)\right] \\
& \quad=t_{0}(2 \Phi(k)-1)^{2}-t_{3}+4 t_{2} \Phi(k)-4 t_{1}(\Phi(k))^{2} . \tag{9}
\end{align*}
$$

It follows that:

$$
\begin{equation*}
w=\Phi^{-1}\left(\frac{-4 B \pm \sqrt{16 B^{2}-16 A C}}{8 A}\right), \tag{10}
\end{equation*}
$$

where:

$$
\begin{aligned}
& A=t_{3}-2 t_{2}+t_{1} \\
& B=-t_{3}+2 t_{2} \Phi(k)+t_{2}-2 t_{1} \Phi(k) \\
& C=-\left[t_{0}(2 \Phi(k)-1)^{2}-t_{3}+4 t_{2} \Phi(k)-4 t_{1}(\Phi(k))^{2}\right]
\end{aligned}
$$

## Performance measurement for the ASI $\boldsymbol{Z}_{\boldsymbol{X}}$ and $\boldsymbol{Z}_{\boldsymbol{e}}$ control charts

The statistical efficiency of a control chart is measured by the speed of detecting process mean shifts. The average run length (ARL) is often used to measure the speed of any FSI control charts. However, the sampling interval of the proposed $Z_{X}$ and $Z_{e}$ control charts in this article is variable, not fixed, so the detecting speed cannot be measured by ARL. Instead, the adjusted average time to signal (AATS) is used to measure the performance of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts. The AATS is the average time from any one of the two process steps' means shift until a signal and all occurred assignable causes are removed. A smaller

AATS is desirable to detect the out-of-control process faster and avoid the losses of defective products.

The average time of a process cycle (ATC) is the average time from the start of production until the signal after the process is checked and adjusted.
Denote $T_{i}$ to be the occurrence time of $\mathrm{AC} i$, which follows an exponential distribution with parameter $\lambda_{i}, i=1$, 2 , where $T_{1}$ and $T_{2}$ are independent.
Denote $T$ to be the occurrence time of the first assignable cause, that is:

$$
T=\min \left(T_{1}, T_{2}\right),
$$

where $T \sim \exp \left(\lambda_{1}+\lambda_{2}\right)$. Hence, the relationship among the AATS, ATC, and the expected $T$ is expressed as:

$$
\begin{equation*}
A A T S=A T C-E(T)=A T C-\frac{1}{\lambda_{1}+\lambda_{2}} \tag{11}
\end{equation*}
$$

Before obtaining the AATS of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts, we have to derive the ATC. The Markov chain approach is used to compute the ATC, due to the nomemory property of the exponential distribution. According to the locations of the sample statistics, 17 possible process states are defined. Table 2 shows the 17 process states. When at least one chart signals, we stop monitoring the process no matter whether the signal is true or false. Hence, the 17 states can be classified into absorbing and transient states. The absorbing states are reached when at least one sample statistic falls out of the control limits; that is, in region $I_{X 4}$ or $I_{e 4}$. The transient states are reached when both sample statistics fall within the control limits. The states from 1 to 16 are defined as the transient states, and state 17 is defined as the absorbing state.

Table 2 The 17 process states

| State | Does AC1 occur? | The location of sample statistic $Z_{X}$ | Is an alarm in the first step? | Does AC2 occur? | The location of sample statistic $Z_{e}$ | Is an alarm in the second step? | Transient state or absorbing state? |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | No | $I_{X 1}$ | No alarm | No | $I_{e 1}$ | No alarm | Transient state |
| 2 | No | $I_{X 1}$ |  | Yes | $I_{e 1}$ |  |  |
| 3 | Yes | $I_{X 1}$ |  | No | $I_{e 1}$ |  |  |
| 4 | Yes | $I_{X 1}$ |  | Yes | $I_{e 1}$ |  |  |
| 5 | No | $I_{X 2}$ |  | No | $I_{e 1}$ |  |  |
| 6 | No | $I_{X 2}$ |  | Yes | $I_{e 1}$ |  |  |
| 7 | Yes | $I_{X 2}$ |  | No | $I_{e 1}$ |  |  |
| 8 | Yes | $I_{X 2}$ |  | Yes | $I_{e 1}$ |  |  |
| 9 | No | $I_{X 1}$ |  | No | $I_{e 2}$ |  |  |
| 10 | No | $I_{X 1}$ |  | Yes | $I_{e 2}$ |  |  |
| 11 | Yes | $I_{X 1}$ |  | No | $I_{e 2}$ |  |  |
| 12 | Yes | $I_{X 1}$ |  | Yes | $I_{e 2}$ |  |  |
| 13 | No | $I_{X 2}$ |  | No | $I_{e 2}$ |  |  |
| 14 | No | $I_{X 2}$ |  | Yes | $I_{e 2}$ |  |  |
| 15 | Yes | $I_{X 2}$ |  | No | $I_{e 2}$ |  |  |
| 16 | Yes | $I_{X 2}$ |  | Yes | $I_{e 2}$ |  |  |
| 17 | At least | one False signal |  | or True | signal |  | Absorbing state |

Denote $\mathbf{P}$ to be the transition probability matrix, and $\mathbf{P}$ is a square matrix of order 17 . Denote $P_{i, j}\left(t_{k}\right)$ to be the transition probability from prior state $i$ to current state $j$ with sampling interval $t_{k}$, where $t_{k}$ is determined by the prior state $i, i=1,2, \ldots, 17$, and $k=1,2,3$. For example, the transition probability from state 1 to state 4 with sampling interval $t_{3}$ is calculated as:

$$
\begin{aligned}
p_{1,4}\left(t_{3}\right)= & P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \\
& \times\left(1-e^{-\lambda_{1} t_{3}}\right) \times\left(1-e^{-\lambda_{2} t_{3}}\right) \\
= & \left(\Phi\left(w_{X}-\delta_{1}\right)-\Phi\left(-w_{X}-\delta_{1}\right)\right) \\
& \times\left(\Phi\left(w_{e}-\delta_{2}\right)-\Phi\left(-w_{e}-\delta_{2}\right)\right) \\
& \times\left(1-e^{-\lambda_{1} t_{3}}\right) \times\left(1-e^{-\lambda_{2} t_{3}}\right)
\end{aligned}
$$

The calculation of all transition probabilities is shown in the Appendix.

Since we consider that the shift in the process mean does not occur at the beginning, hence, the in-control process state would be on $1,5,9$, or 13 . The states $1,5,9$, and 13 are transient states, but state 17 is an absorbing state. Based on the definition of ATC, the ATC is calculated as the sum of (the expected time to reach an absorbing state from the transient state $i) \times($ the corresponding starting probability for the transient state $i$ ), $i=1,5,9,13$; that is:
$A T C=b_{1} M_{1}+b_{5} M_{5}+b_{9} M_{9}+b_{13} M_{13}$,
where $b_{i}$ is the starting probability for the transient state $i$, $i=1,5,9,13$ and $M_{i}$ is the expected time to reach any absorbing state from the transient state $i, i=1,5,9,13$.

The $M_{i}, i=1,2,3, \ldots, 16$, can be derived using the property of Markov chain; that is:

$$
\begin{equation*}
\left(M_{1}, M_{2}, M_{3}, \ldots, M_{16}\right)^{\prime}=(\mathbf{I}-\mathbf{Q})^{-1} \mathbf{t} \tag{13}
\end{equation*}
$$

where $\mathbf{I}$ is the identity matrix of order $16 ; \mathbf{Q}$ is the transition probability matrix, where each element represents the transition probability, $P_{i, j}\left(t_{k}\right)$, from transient state $i$ to transient state $j, i=1, \ldots, 16, j=1, \ldots, 16$; and $\mathbf{t}^{\prime}=\left(t_{3}, t_{3}, t_{3}, t_{3}, t_{2}\right.$, $\left.t_{2}, t_{2}, t_{2}, t_{2}, t_{2}, t_{2}, t_{2}, t_{1}, t_{1}, t_{1}, t_{1}\right)$ is the vector of the sampling intervals from state 1 to state 16 .

Consequently, the AATS is obtained using Eq. 11.

## An example

In this section, we illustrate how to construct and use the proposed ASI $Z_{X}$ and $Z_{e}$ control charts through an empirical automobile braking system example. The data of this example are the measurements of a component part from the automobile braking system. The variables $X=$ ROLLWT was measured for the first step and $Y=$ BAKEWT was measured for the second step. For every sampling, a sample of size 1 is taken at the end of the second step, and a pair of observations $(X, Y)$ is measured. Two machines are
operated in the two dependent process steps. The time $\left(T_{1}\right)$ until machine 1 is out of control follows an exponential distribution with parameter $\lambda_{1}=0.03$, and the time $\left(T_{2}\right)$ until machine 2 is out of control follows an exponential distribution with parameter $\lambda_{2}=0.04$. The variables $T_{1}$ and $T_{2}$ are independent. From the process history, the means of the process may be shifted when two machines are out of control. The out-of-control machine 1 only influences the mean of $X$ on the first process step and the standard deviation is unaffected. The out-of-control machine 2 only influences the mean of $Y$ on the second process step and the standard deviation is unaffected. We collect 45 pairs of observations $(X, Y)$, from the in-control process, to establish their statistical relationship.

From the results of data analysis, both variables $X$ and $Y$ follow bivariate normal distributions. Variable $Y$ is affected by variable $X$, and their relationship is expressed by a simple linear regression model. Using the least square method for the 45 observations, the fitted regression model is:

$$
\begin{equation*}
\widehat{Y} \mid X=93.2+0.513 X \tag{14}
\end{equation*}
$$

From the model assumptions checking, the fitted regression model is appropriated. So, we use Eq. 14 to calculate the estimated values $(\widehat{Y} \mid X)$ and the residuals (e). The estimated means and standard deviations of variables $X$ and $e$ are $\widehat{\mu}_{X}=210.25, \widehat{\sigma}_{X}=1.19, \widehat{\mu}_{e}=0$, and $\widehat{\sigma}_{e}=0.88$. Hence, when both steps are in control, $X \sim \mathrm{~N}\left(210.25,1.19^{2}\right)$ and $e \sim \mathrm{~N}$ $\left(0,0.88^{2}\right)$. According to the historical data, the shift scale of the process mean caused by the out-of-control machine 1 is $\delta_{1} \sigma_{X}$, where $\delta_{1}=0.5$, and the shift scale of the process mean caused by the out-of-control machine 2 is $\delta_{2} \sigma_{e}$, where $\delta_{2}=0.75$. Hence, the distribution of $X$ becomes $X \sim \mathrm{~N}(210.25$ $+0.5 \times 1.19,1.19^{2}$ ) for the out-of-control first process step, and the distribution of $e$ becomes $e \sim \mathrm{~N}\left(0.75 \times 0.88,0.88^{2}\right)$ for the out-of-control second process step.

For three different adaptive sampling intervals, the process engineers determine $t_{1}=0.01 \mathrm{~h}, t_{2}=0.5 \mathrm{~h}$, and $t_{3}=1.15 \mathrm{~h}$; thus, the warning limit is calculated to be 1.58 by using Eq. 10. Consequently, the structures of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts are as follows:

$$
\begin{align*}
& U C L_{Z_{X}}=3 \quad U C L_{Z_{e}}=3  \tag{15}\\
& U W L_{Z_{X}}=1.58 \quad U W L_{Z_{e}}=1.58 \\
& C L_{Z_{X}}=0 \quad C L_{Z_{e}}=0 \\
& L W L_{Z_{X}}=-1.58 \quad L W L_{Z_{e}}=-1.58 \\
& L C L_{Z_{X}}=-3 \quad L C L_{Z_{e}}=-3
\end{align*}
$$

To detect the changes in the process means faster, the constructed ASI $Z_{X}$ and $Z_{e}$ control charts are used. The first sampling interval is randomly chosen to be 0.5 h . After 0.5 $h$, the first sample is taken from the end of the second process step, and the paired observation of $(X, Y)$ is (209, 201). Hence, the calculated values of $(X, e)$ is $(209,0.583)$
by using Eqs. 2 and 14. Using Eqs. 2 and 3, the values of $Z_{X}$ and $Z_{e}$ are as follows:
$Z_{X}=\frac{209-210.25}{1.19}=-1.05$
$Z_{e}=\frac{0.583}{0.88}=0.66$
Both sample statistics $(-1.05,0.66)$ are within the warning limits, so we adopt the longer sample interval $t_{3}=1.15 \mathrm{~h}$. After 1.15 h , the second sample is taken, and the paired observation of $(X, Y)$ is $(208,202)$. Hence, the sample statistics $\left(Z_{X}, Z_{e}\right)$ of the second sample are $(-1.89,2.38)$. Both of these statistics fall outside the warning limits, but inside the control limits, so the shorter sampling interval $t_{1}=0.01 \mathrm{~h}$ is adopted; that is, the next sample is taken 0.01 h after the second sample. Once at least one of the sample statistics falls outside the control limits, then the process is stopped and the out-of-control machine 1 or/and machine 2 needs to be repaired.

The AATS is used to measure the performance of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts. There are 17 possible process states for the empirical example, which are the same as the process states presented in Table 2. The ATC is calculated to be 62.80 h by using Eq. 12. Hence, the AATS is 52.51 h by using Eq. 11. It means that the average time to detect the out-of-control process is 52.51 h .

For comparing the AATS of the ASI $Z_{X}$ and $Z_{e}$ control charts with that of the FSI $Z_{X}$ and $Z_{e}$ control charts, we also calculate the AATS of the FSI $Z_{X}$ and $Z_{e}$ control charts. The approach to obtain the ATC and AATS of the FSI $Z_{X}$ and $Z_{e}$ control charts is similar to that of the ASI $Z_{X}$ and $Z_{e}$ control charts, except that there are no warning limits on the FSI $Z_{X}$ and $Z_{e}$ control charts. After calculation, the AATS of the FSI $Z_{X}$ and $Z_{e}$ control charts is 54.91 h . It shows that the proposed ASI $Z_{X}$ and $Z_{e}$ control charts can detect the shift of the process means faster than the FSI $Z_{X}$ and $Z_{e}$ control charts by about 2 h . Hence, the proposed ASI $Z_{X}$ and $Z_{e}$ control charts offer better performance than the FSI $Z_{X}$ and $Z_{e}$ control charts.

## Comparisons with FSI $Z_{X}$ and $Z_{e}$ control charts

Following the example in Sect. 6, we survey the performance of the proposed charts and the FSI charts for various combinations of shift scales $\left(\delta_{1}, \delta_{2}\right)$ of the process means and parameters $\left(\lambda_{1}, \lambda_{2}\right)$. According to the historical data, the shift scales $\left(\delta_{1}, \delta_{2}\right)$ of the process means caused by the out-of-control machines 1 and 2 range from 0.5 to 1.5 , and the parameters $\left(\lambda_{1}, \lambda_{2}\right)$ range from 0.03 to 0.05 , respectively. The AATS under various combinations of $\left(\delta_{1}, \delta_{2}\right)$ and $\left(\lambda_{1}, \lambda_{2}\right)$ are calculated and illustrated in Table 3 and Table 4, respectively.

The numerical results in Tables 3 and 4 show that, when detecting small and moderate shifts in process means, the performance of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts is much better than the FSI $Z_{X}$ and $Z_{e}$ control charts. Using ASI $Z_{X}$ and $Z_{e}$ control charts instead of FSI $Z_{X}$ and $Z_{e}$ control charts can reduce the AATS to about 0.5 h to 5.5 h . In addition, the proposed ASI $Z_{X}$ and $Z_{e}$ control charts have better performance in detecting small shifts ( $\delta_{1} \leq 1$ and $\delta_{2} \leq 1$ ) than moderate shifts $\left(1<\delta_{1} \leq 1.5\right.$ or $\left.1<\delta_{2} \leq 1.5\right)$. Under small shifts ( $\delta_{1} \leq 1$ and $\delta_{2} \leq 1$ ) in the process means, the ASI $Z_{X}$ and $Z_{e}$ control charts can decrease the AATS to about 3.5 h to 5.5 h , compared to the FSI $Z_{X}$ and $Z_{e}$ control charts. Under moderate shifts $\left(1<\delta_{1} \leq 1.5\right.$ or $\left.1<\delta_{2} \leq 1.5\right)$ in process means, the ASI $Z_{X}$ and $Z_{e}$ control charts can decrease the AATS to about 0.3 h to 3.8 h , compared to the FSI $Z_{X}$ and $Z_{e}$ control charts. Furthermore, the large values of $\lambda_{1}$ and $\lambda_{2}$ lead to a smaller AATS than the small values of $\lambda_{1}$ and $\lambda_{2}$.

One of the advantages of Shewhart control charts is their speed in detecting large shifts in the process mean. However, it is not efficient to use Shewhart control charts when detecting small and moderate shifts in process means. The proposed ASI $Z_{X}$ and $Z_{e}$ control charts may improve this problem for two dependent process steps. Based on the numerical results, the performance of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts always outperform the FSI $Z_{X}$ and $Z_{e}$ control charts.

## Conclusions

In this article, the adaptive sampling interval (ASI) $Z_{X}$ and $Z_{e}$ cause-selecting control charts are proposed to effectively detect and distinguish which one of two dependent process steps is out of control.

The adjusted average time to signal (AATS) is calculated to measure the performance of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts by the Markov chain approach. An empirical automobile braking system example illustrates the application of the proposed ASI $Z_{X}$ and $Z_{e}$ control charts and shows that the performance of the ASI $Z_{X}$ and $Z_{e}$ control charts is better than the FSI $Z_{X}$ and $Z_{e}$ control charts. From some numerical results, it is demonstrated that the proposed ASI $Z_{X}$ and $Z_{e}$ control charts outperform the FSI $Z_{X}$ and $Z_{e}$ control charts. It implicates that a process engineer can identify an out-of-control process step quickly when using the proposed ASI $Z_{X}$ and $Z_{e}$ control charts, and it also leads a reduction in the losses of defective products.

Several important extensions of the developed model can be expanded. It is straightforward to extend the proposed model to study adaptive sample size (ASS), adaptive sample size and sampling interval (ASSI) cause-selecting control charts, or other control charts, like adaptive cause-selecting control charts in the presence of correlation, exponentially weighted moving average (EWMA), or cumulative sum (CUSUM) cause-selecting control charts. One particularly interesting research area for future research involves the
Table 3 Comparison of the AATS of the ASI and FSI $Z_{X}$ and $Z_{e}$ control charts when $\lambda_{1}=0.03$ and $\lambda_{2}=0.04$

|  |  |  |  | $\delta_{1}$ | 0.5 |  |  |  |  | 0.75 |  |  |  |  | 1 |  |  |  |  | 1.25 |  |  |  |  | 1.5 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $t_{1}$ | $t_{2}$ | $t_{3}$ | w | $\delta_{2}$ | 0.5 | 0.75 | 1 | 1.25 | 1.5 | 0.5 | 0.75 | 1 | 1.25 | 1.5 | 0.5 | 0.75 | 1 | 1.25 | 1.5 | 0.5 | 0.75 | 1 | 1.25 | 1.5 | 0.5 | 0.75 | 1 | 1.25 | 1.5 |
| 0.01 | 0.1 | 1.15 | 1.7887 |  | 74.6527 | 51.7056 | 34.2193 | 23.0980 | 16.6449 | 53.6831 | 40.7301 | 28.9235 | 20.3199 | 14.8787 | 37.6643 | 30.6475 | 23.1768 | 16.9352 | 12.5834 | 27.4356 | 23.2154 | 18.2509 | 13.6685 | 10.2052 | 21.4728 | 18.4556 | 14.7246 | 11.0883 | 8.2014 |
| 0.01 | 0.5 | 1.15 | 1.5754 |  | 75.3654 | 52.5110 | 35.0544 | 23.8975 | 17.3649 | 54.4715 | 41.5253 | 29.7116 | 21.0710 | 15.5610 | 38.4694 | 31.4171 | 23.9189 | 17.6440 | 13.2374 | 28.2000 | 23.9372 | 18.9458 | 14.3425 | 10.8422 | 22.1592 | 19.1079 | 15.3622 | 11.7226 | 8.8186 |
| 0.09 | 0.1 | 1.15 | 1.7874 |  | 74.6722 | 51.7294 | 34.2430 | 23.1185 | 16.6614 | 53.7069 | 40.7588 | 28.9523 | 20.3451 | 14.8990 | 37.6879 | 30.6765 | 23.2073 | 16.9630 | 12.6063 | 27.4562 | 23.2411 | 18.2792 | 13.6955 | 10.2284 | 21.4894 | 18.4765 | 14.7484 | 11.1121 | 8.2226 |
| 0.09 | 0.5 | 1.15 | 1.5718 |  | 75.4029 | 52.5563 | 35.0983 | 23.9346 | 17.3936 | 54.5164 | 41.5779 | 29.7632 | 21.1151 | 15.5955 | 38.5130 | 31.4689 | 23.9721 | 17.6915 | 13.2757 | 28.2370 | 23.9821 | 18.9941 | 14.3879 | 10.8805 | 22.1880 | 19.1434 | 15.4020 | 11.7617 | 8.8531 |
| 0.01 | 0.1 | 1.5 | 1.2917 |  | 73.5385 | 50.4976 | 33.4215 | 22.9524 | 17.1516 | 52.5264 | 39.6559 | 28.2447 | 20.2256 | 15.3834 | 36.9346 | 29.9955 | 22.7865 | 16.9726 | 13.1086 | 27.3462 | 23.1394 | 18.2764 | 13.9347 | 10.8161 | 22.0105 | 18.9481 | 15.1966 | 11.6497 | 8.9795 |
| 0.01 | 0.5 | 1.5 | 1.1095 |  | 74.9397 | 52.0051 | 34.8842 | 24.2579 | 18.2485 | 54.040 | 41.0976 | 29.6016 | 21.4430 | 16.4228 | 38.3483 | 31.3240 | 24.0292 | 18.1104 | 14.1087 | 28.5989 | 24.3167 | 19.3992 | 15.0005 | 11.7903 | 23.0625 | 19.9537 | 16.1881 | 12.6314 | 9.9126 |
| 9 | 0.1 | 1.5 | 1.2878 |  | 73.6228 | 50.5955 | 33.5122 | 23.0448 | 17.2033 | 52.6231 | 39.7631 | 28.3450 | 20.3071 | 15.4428 | 37.0240 | 30.0960 | 22.8848 | 17.0561 | 13.1720 | 27.4177 | 23.2216 | 18.3610 | 14.0109 | 10.8771 | 22.0618 | 19.0089 | 15.2625 | 11.7127 | 9.0324 |
| 0.09 | 0.5 | 1.5 | 1.1018 |  | 75.0603 | 52.1407 | 35.0074 | 24.3526 | 18.3117 | 54.1378 | 41.2434 | 29.7753 | 21.5482 | 16.4952 | 38.4693 | 31.4573 | 24.1572 | 18.2165 | 14.1856 | 28.6921 | 24.4223 | 19.5066 | 15.0954 | 11.8633 | 23.1249 | 20.0278 | 16.2683 | 12.7071 | 9.9741 |
|  |  |  |  |  | 77.7839 | 55.5624 | 38.0965 | 26.4924 | 19.3611 | 55.5624 | 43.6206 | 32.2752 | 23.5336 | 17.6068 | 38.0965 | 32.2752 | 25.7009 | 19.7653 | 15.2274 | 30.0017 | 25.8075 | 20.7921 | 16.0022 | 12.1730 | 23.4402 | 20.4703 | 16.7788 | 13.0804 | 9.9750 |
|  | rence |  |  |  | 3.79 | 4.41 | 3.95 | 2.84 | 2.07 | 4.28 | 4.37 | 3.80 | 2.76 | 2.08 | 3.76 | 3.72 | 3.31 | 2.56 | 2.06 | 2.66 | 2.67 | 2.54 | 2.33 | 1.97 | 1.97 | 2.01 | 2.05 | 1.99 | 1.77 |
|  | rence 2 |  |  |  | 1.93 | 2.35 | 2.27 | 1.44 | 0.41 | 2.29 | 2.44 | 2.29 | 1.44 | 0.46 | 2.18 | 2.25 | 1.94 | 1.28 | 0.46 | 1.31 | 1.39 | 1.29 | 0.91 | 0.31 | 0.32 | 0.44 | 0.51 | 0.37 | 0.00 | Difference $1=$ AATS of FSI control chart-max. AATS of ASI control chart

Difference $2=$ AATS of FSI control chart-min. AATS of ASI control chart

Table 4 Comparison of the AATS of the ASI and FSI $Z_{X}$ and $Z_{e}$ control charts when $\lambda_{1}=0.05$ and $\lambda_{2}=0.05$

|  |  |  |  | $\delta_{1}$ | 0.5 |  |  |  |  | 0.75 |  |  |  |  | 1 |  |  |  |  | 1.25 |  |  |  |  | 1.5 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $t_{2}$ | $t_{3}$ | $w$ | $\delta_{2}$ | 0.5 | 0.75 | 1 | 1.25 | 1.5 | 0.5 | 0.75 | 1 | 1.25 | 1.5 | 0.5 | 0.75 | 1 | 1.25 | 1.5 | 0.5 | 0.7 | 1 | 1.25 | 1.5 | 0.5 | 0.75 | 1 | 1.25 | 1.5 |
| 0.01 | 0.1 | 1.15 | 1.7887 |  | 74.6350 | 51.9107 | 34.5376 | 23.4292 | 16.9429 | 51.9107 | 39.8411 | 28.7234 | 20.4930 | 15.1955 | 34.5376 | 28.7234 | 22.3663 | 16.8552 | 12.8601 | 23.4292 | 20.4930 | 16.8552 | 13.2597 | 10.3462 | 16.9429 | 15.1955 | 12.8601 | 10.3462 | 8.1408 |
| 0.01 | 0.5 | 1.15 | 1.5754 |  | 75.3748 | 52.7372 | 35.3853 | 24.2351 | 17.6661 | 52.7372 | 40.6619 | 29.5199 | 21.2404 | 15.8695 | 35.3853 | 29.5200 | 23.1114 | 17.5485 | 13.4922 | 24.2351 | 21.2404 | 17.5485 | 13.9092 | 10.9509 | 17.6661 | 15.8695 | 13.4922 | 10.9509 | 8.7206 |
| 0.09 | 0.1 | 1.15 | 1.7874 |  | 74.6567 | 51.9371 | 34.5640 | 23.4523 | 16.9616 | 51.9371 | 39.8732 | 28.7561 | 20.5221 | 15.2190 | 34.5640 | 28.7561 | 22.4012 | 16.8877 | 12.8873 | 23.4523 | 20.5221 | 16.8877 | 13.2914 | 10.3739 | 16.9616 | 15.2191 | 12.8873 | 10.3739 | 8.1662 |
| 9 | 0.5 | 1.15 | 1.5718 |  | 75.4170 | 52.7876 | 35.4346 | 24.2772 | 17.6992 | 52.7876 | 40.7209 | 29.5786 | 21.2913 | 15.9099 | 35.4346 | 29.5786 | 23.1724 | 17.6040 | 13.5377 | 24.2772 | 21.2913 | 17.6040 | 13.9624 | 10.9967 | 17.6992 | 15.9099 | 13.5377 | 10.9967 | 8.7619 |
| 0.01 | 0.1 | 1.5 | 1.2917 |  | 72.7925 | 50.0009 | 33.0790 | 22.6582 | 16.8475 | 50.0009 | 38.0553 | 27.3907 | 19.7887 | 15.1138 | 33.0790 | 27.3907 | 21.3589 | 16.3168 | 12.8237 | 22.6582 | 19.7887 | 16.3168 | 12.9931 | 10.4202 | 16.8475 | 15.1138 | 12.8237 | 10.4202 | 8.4074 |
| 0.01 | 0.5 | 1.5 | 1.1095 |  | 74.2286 | 51.5284 | 34.5424 | 23.9508 | 17.9256 | 51.5284 | 39.5195 | 28.7355 | 20.9718 | 16.1128 | 34.5424 | 28.7355 | 22.5722 | 17.3956 | 13.7592 | 23.9508 | 20.9718 | 17.3956 | 13.9818 | 11.3138 | 17.9256 | 16.1128 | 13.7592 | 11.3138 | 9.2532 |
| 0.09 | 0.1 | 1.5 | 1.2878 |  | 72.8904 | 50.1130 | 33.1835 | 22.7432 | 16.9104 | 50.1130 | 38.1791 | 27.5076 | 19.8853 | 15.1863 | 33.1835 | 27.5076 | 21.4741 | 16.4164 | 12.9014 | 22.7431 | 19.8853 | 16.4164 | 13.0841 | 10.4950 | 169104 | 15.1863 | 12.9013 | 10.4950 | 8.4724 |
| 0.09 | 0.5 | 1.5 | 1.1018 |  | 74.3708 | 51.6861 | 34.6868 | 24.0651 | 18.0065 | 51.6861 | 39.6903 | 28.8934 | 21.0992 | 16.2049 | 34.6868 | 28.8934 | 22.7248 | 17.5248 | 13.8569 | 24.0651 | 21.0992 | 17.5248 | 14.0979 | 11.4070 | 18.0065 | 16.2049 | 13.8569 | 11.4070 | 9.3324 |
| FSI control chats |  |  |  |  | 77.3323 | 54.9085 | 37.3681 | 25.7887 | 18.7195 | 56.8085 | 44.0212 | 32.0490 | 22.9898 | 16.9562 | 40.6935 | 33.7139 | 26.1005 | 19.4987 | 14.6413 | 26.4924 | 23.5336 | 19.7653 | 15.8982 | 12.5911 | 19.3611 | 17.6068 | 15.2274 | 12.5911 | 10.1543 |
| Difference 1 |  |  |  |  | 4.99 | 5.56 | 5.02 | 3.83 | 2.51 | 5.56 | 5.57 | 4.88 | 3.74 | 2.49 | 5.02 | 4.88 | 4.34 | 3.45 | 2.40 | 3.83 | 3.74 | 3.45 | 2.91 | 2.24 | 2.51 | 2.49 | 2.40 | 2.24 | 2.01 |
| Difference 2 |  |  |  |  | 2.37 | 2.77 | 2.66 | 2.22 | 1.35 | 2.77 | 2.90 | 2.70 | 2.24 | 1.40 | 2.66 | 2.70 | 2.53 | 2.16 | 1.37 | 2.22 | 2.24 | 2.16 | 1.80 | 1.18 | 1.35 | 1.40 | 1.37 | 1.18 | 0.82 |

Difference $1=$ AATS of FSI control chart-max. AATS of ASI control chart
Difference $2=$ AATS of FSI control chart-min. AATS of ASI control chart
adaptive economic statistical modeling of dependent process steps subject to multiple assignable causes.

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## Appendix

The calculation of the transition probabilities are as follows:

$$
\begin{aligned}
& P_{i j}\left(t_{k}\right) \quad i=1, \ldots, 17, j=1, \ldots, 17, k=1,2,3 \\
& P_{1,1}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X}\right] \times P\left[\left|Z_{e}\right|<w_{e}\right] \times e^{-\lambda_{1} t_{3}} \times e^{-\lambda_{2} t_{3}} \\
& P_{1,2}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \times e^{-\lambda_{1} t_{3}} \times\left(1-e^{-\lambda_{2} t_{3}}\right) \\
& P_{1,3}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right) \times e^{-\lambda_{2} t_{3}} \\
& P_{1,4}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right) \times\left(1-e^{-\lambda_{2} t_{3}}\right) \\
& P_{1,5}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X}\right] \times P\left[\left|Z_{e}\right|<w_{e}\right] \times e^{-\lambda_{1} t_{3}} \times e^{-\lambda_{2} t_{3}} \\
& P_{1,6}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \times e^{-\lambda_{1} t_{3}} \times\left(1-e^{-\lambda_{2} t_{3}}\right) \\
& P_{1,7}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right) \times e^{-\lambda_{2} t_{3}} \\
& P_{1,8}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right) \times\left(1-e^{-\lambda_{2} t_{3}}\right) \\
& P_{1,9}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e}\right] \times e^{-\lambda_{1} t_{3}} \times e^{-\lambda_{2} t_{3}} \\
& P_{1,10}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right] \times e^{-\lambda_{1} t_{3}} \times\left(1-e^{-\lambda_{2} t_{3}}\right) \\
& P_{1,11}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right) \times e^{-\lambda_{2} t_{3}} \\
& P_{1,12}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right) \times\left(1-e^{-\lambda_{2} t_{3}}\right) \\
& P_{1,13}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e}\right] \times e^{-\lambda_{1} t_{3}} \times e^{-\lambda_{2} t_{3}} \\
& P_{1,14}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e e} \mid \delta_{2}\right] \times e^{-\lambda_{1} t_{1}} \times\left(1-e^{-\lambda_{2} t_{1}}\right) \\
& P_{1,15}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right) \times e^{-\lambda_{2} t_{3}} \\
& P_{1,16}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right] \\
& \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right) \times\left(1-e^{-\lambda_{2} t_{3}}\right) \\
& P_{1,17}\left(t_{3}\right)=1-\sum_{j=1}^{16} P_{1, j}\left(t_{3}\right)
\end{aligned}
$$

$$
\begin{aligned}
& P_{2,1}\left(t_{3}\right)=0 \\
& P_{2,2}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \times e^{-\lambda_{1} t_{3}} \\
& P_{2,3}\left(t_{3}\right)=0 \\
& P_{2,4}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right) \\
& P_{2,5}\left(t_{3}\right)=0 \\
& P_{2,6}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \times e^{-\lambda_{1} t_{3}} \\
& P_{2,7}\left(t_{3}\right)=0 \\
& P_{2,8}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right)
\end{aligned}
$$

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\(P_{2,9}\left(t_{3}\right)=0\)
\(P_{2,10}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right] \times e^{-\lambda_{1} t_{3}}\)
\(P_{2,11}\left(t_{3}\right)=0\)
\(P_{2,12}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right)\)
\(P_{2,13}\left(t_{3}\right)=0\)
\(P_{2,14}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right] \times e^{-\lambda_{1} t_{3}}\)
\(P_{2,15}\left(t_{3}\right)=0\)
\(P_{2,16}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{1} t_{3}}\right)\)
\(P_{2,17}\left(t_{3}\right)=1-\sum_{j=1}^{16} P_{2, j}\left(t_{3}\right)\)
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$P_{3,1}\left(t_{3}\right)=0$
$P_{3,2}\left(t_{3}\right)=0$
$P_{3,3}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e}\right] \times e^{-\lambda_{2} t_{3}}$
$P_{3,4}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{2} t_{3}}\right)$
$P_{3,5}\left(t_{3}\right)=0$
$P_{3,6}\left(t_{3}\right)=0$
$P_{3,7}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e}\right] \times e^{-\lambda_{2} t_{3}}$
$P_{3,8}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right] \times P\left[\left|Z_{e}\right|<w_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{2} t_{3}}\right)$
$P_{3,9}\left(t_{3}\right)=0$
$P_{3,10}\left(t_{3}\right)=0$
$P_{3,11}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e}\right] \times e^{-\lambda_{2} t_{3}}$
$P_{3,12}\left(t_{3}\right)=P\left[\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{2} t_{3}}\right)$
$P_{3,13}\left(t_{3}\right)=0$
$P_{3,14}\left(t_{3}\right)=0$
$P_{3,15}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e}\right] \times e^{-\lambda_{2} t_{3}}$
$P_{3,16}\left(t_{3}\right)=P\left[-k_{X}<Z_{X} \leq-w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right] \times P\left[-k_{e}<Z_{e} \leq-w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right] \times\left(1-e^{-\lambda_{2} t_{3}}\right)$
$P_{3,17}\left(t_{3}\right)=1-\sum_{j=1}^{16} P_{3, j}\left(t_{3}\right)$
$P_{4,1}\left(t_{3}\right)=P_{4,2}\left(t_{3}\right)=P_{4,3}\left(t_{3}\right)=0$
$P_{4,4}\left(t_{3}\right)=P\left(\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right) \times P\left(Z_{e}<w_{e} \mid \delta_{2}\right)$
$P_{4,5}\left(t_{3}\right)=P_{4,6}\left(t_{3}\right)=P_{4,7}\left(t_{3}\right)=0$
$P_{4,8}\left(t_{3}\right)=P\left(-k_{X}<Z_{X} \leq w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right) \times P\left(Z_{e}<w_{e} \mid \delta_{2}\right)$
$P_{4,9}\left(t_{3}\right)=P_{4,10}\left(t_{3}\right)=P_{4,11}\left(t_{3}\right)=0$
$P_{4,12}\left(t_{3}\right)=P\left(\left|Z_{X}\right|<w_{X} \mid \delta_{1}\right) \times P\left(-k_{e}<Z_{e} \leq w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right)$
$P_{4,13}\left(t_{3}\right)=P_{4,14}\left(t_{3}\right)=P_{4,15}\left(t_{3}\right)=0$
$P_{4,16}\left(t_{3}\right)=P\left(-k_{X}<Z_{X} \leq w_{X} \cup w_{X} \leq Z_{X}<k_{X} \mid \delta_{1}\right) \times P\left(-k_{e}<Z_{e} \leq w_{e} \cup w_{e} \leq Z_{e}<k_{e} \mid \delta_{2}\right)$
$P_{4,17}\left(t_{3}\right)=1-\sum_{j=1}^{16} P_{4, j}\left(t_{3}\right)$

The transition probabilities for $P_{5, j}\left(t_{2}\right), P_{6, j}\left(t_{2}\right), P_{7, j}\left(t_{2}\right)$, $P_{8, j}\left(t_{2}\right), P_{9, j}\left(t_{2}\right), P_{10, j}\left(t_{2}\right), P_{11, j}\left(t_{2}\right)$, and $P_{12, j}\left(t_{2}\right)$ are calculated by replacing $t_{2}$ on $t_{3}$ for $P_{1, j}\left(t_{3}\right), P_{2, j}\left(t_{3}\right), P_{3, j}\left(t_{3}\right)$, and $P_{4, j}\left(t_{3}\right), j=1,2, \ldots, 17$ :

$$
\begin{array}{ll}
P_{5, j}\left(t_{2}\right)=P_{1, j}\left(t_{2}\right) & j=1,2, \ldots, 17 \\
P_{6, j}\left(t_{2}\right)=P_{2, j}\left(t_{2}\right) & j=1,2, \ldots, 17 \\
P_{7, j}\left(t_{2}\right)=P_{3, j}\left(t_{2}\right) & j=1,2, \ldots, 17 \\
P_{8, j}\left(t_{2}\right)=P_{4, j}\left(t_{2}\right) & j=1,2, \ldots, 17 \\
P_{9, j}\left(t_{2}\right)=P_{1, j}\left(t_{2}\right) & j=1,2, \ldots, 17 \\
P_{10, j}\left(t_{2}\right)=P_{2, j}\left(t_{2}\right) & j=1,2, \ldots, 17 \\
P_{11, j}\left(t_{2}\right)=P_{3, j}\left(t_{2}\right) & j=1,2, \ldots, 17 \\
P_{12, j}\left(t_{2}\right)=P_{4, j}\left(t_{2}\right) & j=1,2, \ldots, 17
\end{array}
$$

The transition probabilities for $P_{13, j}\left(t_{1}\right), P_{14, j}\left(t_{1}\right), P_{15, j}\left(t_{1}\right)$, and $P_{16, j}\left(t_{1}\right)$ are calculated by replacing $t_{1}$ on $t_{3}$ for $P_{1, j}\left(t_{3}\right)$, $P_{2, j}\left(t_{3}\right), P_{3, j}\left(t_{3}\right)$, and $P_{4, j}\left(t_{3}\right), j=1,2, \ldots, 17$ :
$P_{13, j}\left(t_{1}\right)=P_{1, j}\left(t_{1}\right) \quad j=1,2, \ldots, 17$
$P_{14, j}\left(t_{1}\right)=P_{2, j}\left(t_{1}\right) \quad j=1,2, \ldots, 17$
$P_{15, j}\left(t_{1}\right)=P_{3, j}\left(t_{1}\right) \quad j=1,2, \ldots, 17$
$P_{16, j}\left(t_{1}\right)=P_{4, j}\left(t_{1}\right) \quad j=1,2, \ldots, 17$
$P_{17, j}=0 j \neq 17$
$P_{17,17}=1$

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